Isoperimetric inequalities in high-dimensional convex sets Boaz Klartag, ETH Zurich 2025

Lecture 6: Log-concave probability measures and isotropic constants

Which probability measures in high dimensions enjoy *concentration phenomena*? With respect to which probability measures on \mathbb{R}^n Lipschitz functions are concentrated near their expectation? For which measures on \mathbb{R}^n most of the mass is located near "any equator", and perhaps even "non-linear equators" which are hypersurfaces partitioning space into two parts of equal mass?

Earlier in these lectures we considered the case of the uniform measure on the sphere unit S^{n-1} , as well as the closely related uniform measure on the Euclidean unit ball B^n . Furthermore, we know that when

$$X = (X_1, \dots, X_n) \sim \text{Unif}(\sqrt{n}S^{n-1})$$

and n is very large while k = o(n), the random variables

$$X_1, \ldots, X_k \in \mathbb{R}^k$$

are approximately independent standard Gaussian random variables in the total variation distance (see Diaconis and Freedman [5, Section 6] for this statement and its history). Thus the standard Gaussian probability measure on \mathbb{R}^n enjoys strong concentration properties, which it inherits from the high-dimensional sphere (see exercise below for a better proof of Gaussian concentration of Lipschitz functions).

There are concentration inequalities available for product measures (i.e., independent random variables), in particular for the boolean cube $\{-1,1\}^n$, and for random variables with weak dependence properties.

Here we study a class of probability measures in \mathbb{R}^n whose concentration properties were understood relatively recently, which are high-dimensional measures with convexity properties. In particular, we focus on log-concave probability measures.

We begin with the Prékopa-Leindler inequality, which is a functional version of the Brunn-Minkowski inequality.

Theorem 6.1 (Prékopa-Leindler). Suppose that $f, g, h : \mathbb{R}^n \to [0, \infty)$ are measurable functions and $0 < \lambda < 1$ are such that for any $x, y \in \mathbb{R}^n$,

$$h\left((1-\lambda)x + \lambda y\right) \ge f(x)^{1-\lambda}g(y)^{\lambda}.\tag{1}$$

Then,

$$\int_{\mathbb{R}^n} h \ge \left(\int_{\mathbb{R}^n} f \right)^{1-\lambda} \left(\int_{\mathbb{R}^n} g \right)^{\lambda}, \tag{2}$$

whenever the integrals on the right-hand side converge.

Remarks.

1. In the case where $A,B\subseteq\mathbb{R}^n$ have finite volume, by setting

$$f = 1_A, \qquad g = 1_B, \qquad h = 1_{(1-\lambda)A + \lambda B}$$

we recover the Brunn-Minkowski inequality in its multiplicative form. Indeed, f, g, h satisfy the requirements of Theorem 6.1, and hence by its conclusion

$$Vol_n\left((1-\lambda)A + \lambda B\right) = \int h$$

$$\geq \left(\int f\right)^{1-\lambda} \left(\int g\right)^{\lambda} = \operatorname{Vol}_n(A)^{1-\lambda} \operatorname{Vol}_n(B)^{\lambda}.$$

There are also several ways to deduce the Prekopa-Leindler inequality from the Brunn-Minkowski inequality. For example, one may consider convex bodies in higher dimensions whose marginal distributions yield the given functions, and apply Brunn-Minkowski (see, e.g., [10]).

2. The Prékopa-Leindler inequality may be viewed as a certain converse to Hölder's inequality. Indeed, the Hölder inequality implies that

$$\int_{\mathbb{R}^n} f^{1-\lambda} g^{\lambda} \le \left(\int_{\mathbb{R}^n} f \right)^{1-\lambda} \left(\int_{\mathbb{R}^n} g \right)^{\lambda}$$

while the Prékopa-Leindler inequality yields

$$\int_{\mathbb{R}^n} \left[\sup_{z=(1-\lambda)y+\lambda z} f(y)^{1-\lambda} g(z)^{\lambda} \right] dx \ge \left(\int_{\mathbb{R}^n} f \right)^{1-\lambda} \left(\int_{\mathbb{R}^n} g \right)^{\lambda}.$$

Proof of Theorem 6.1 for n=1. Consider first the case where f and g are bounded functions. If f or g vanish almost everywhere, then there is nothing to prove. Hence we may assume that $\|f\|_{\infty}$ and $\|g\|_{\infty}$ are positive numbers. In fact, by homogeneity we may assume that

$$||f||_{\infty} = ||g||_{\infty} = 1,$$
 (3)

since otherwise we may replace f by $f/\|f\|_{\infty}$, replace g by $g/\|g\|_{\infty}$ and replace h by $h/\left(\|f\|_{\infty}^{1-\lambda}\|g\|_{\infty}^{\lambda}\right)$, without affecting the validity of neither the assumptions nor the conclusions of the theorem.

Recall that we abbreviate $\{h > t\} = \{x \in \mathbb{R} ; h(x) > t\}$. Observe that condition (1) imply that for all t > 0,

$$\{h > t\} \supseteq (1 - \lambda)\{f > t\} + \lambda\{g > t\}.$$
 (4)

If 0 < t < 1 then both sets on the right-hand side of (4) are non-empty. Hence, by the one-dimensional Brunn-Minkowski inequality (which is a triviality), for 0 < t < 1,

$$m(\{h > t\}) \ge (1 - \lambda)m(\{f > t\}) + \lambda m(\{g > t\}),$$

where m is the Lebesgue measure on the real line. Therefore,

$$\begin{split} \int_{\mathbb{R}} h &= \int_0^\infty m(\{h > t\}) dt \ge \int_0^1 m(\{h > t\}) dt \\ &\ge (1 - \lambda) \int_0^1 m(\{f > t\}) dt + \lambda \int_0^1 m(\{g > t\}) dt \\ &= (1 - \lambda) \int_{\mathbb{R}} f + \lambda \int_{\mathbb{R}} g \ge \left(\int_{\mathbb{R}} f\right)^{1 - \lambda} \left(\int_{\mathbb{R}} g\right)^{\lambda}. \end{split}$$

This concludes the proof in the case where f and g are bounded. For the general case, for M>0 we replace f by $\min\{f,M\}$, we replace g by $\min\{g,M\}$ and h by $\min\{g,M\}$. Such a truncation still satisfies the requirements of the Prékopa-Leindler inequality (with the same function h). Hence, by the case of the inequality that was already proven,

$$\int_{\mathbb{R}} h \geq \left(\int_{\mathbb{R}} \min\{f,M\}\right)^{1-\lambda} \left(\int_{\mathbb{R}} \min\{g,M\}\right)^{\lambda} \xrightarrow{M \to \infty} \left(\int_{\mathbb{R}} f\right)^{1-\lambda} \left(\int_{\mathbb{R}} g\right)^{\lambda},$$

where we used the monotone convergence theorem in the last passage. \Box

Proof of Theorem 6.1 for $n \geq 2$. By induction on n. We use $x = (y, t) \in \mathbb{R}^{n-1} \times \mathbb{R}$ as coordinates in \mathbb{R}^n and set

$$F(y) = \int_{-\infty}^{\infty} f(y,t) dt = \int_{-\infty}^{\infty} f_y(t) dt,$$

$$G(y) = \int_{-\infty}^{\infty} g(y,t) dt = \int_{-\infty}^{\infty} g_y(t) dt,$$

$$H(y) = \int_{-\infty}^{\infty} h(y,t) dt = \int_{-\infty}^{\infty} h_y(t) dt.$$

We claim that if $y = (1 - \lambda)y_1 + \lambda y_2$, for $y_1, y_2 \in \mathbb{R}^n$, then,

$$H(y) \ge F(y_1)^{1-\lambda} G(y_2)^{\lambda}. \tag{5}$$

Indeed, if $t = (1 - \lambda)t_1 + \lambda t_2$ for $t_1, t_2 \in \mathbb{R}$, then

$$h_y(t) \ge f_{y_1}(t_1)^{1-\lambda} f_{y_2}(t_2)^{\lambda}.$$

Hence (5) follows by the one-dimensional Prékopa-Leindler inequality. Thanks to (5) and the induction hypothesis, we may apply the (n-1)-dimensional Prékopa-Leindler inequality for the functions F,G and H and conclude (2). $\hfill\Box$

Definition 6.2. A function $\rho : \mathbb{R}^n \to [0, \infty)$ is log-concave if for all $x, y \in \mathbb{R}^n$ and $0 < \lambda < 1$,

$$\rho\left((1-\lambda)x + \lambda y\right) \ge \rho(x)^{1-\lambda}\rho(y)^{\lambda},$$

i.e., if the set $\Omega = \{\rho > 0\}$ is convex and $-\log \rho$ is a convex function on Ω .

We say that a probability measure (or a random vector) in \mathbb{R}^n is log-concave if it is supported in an affine subspace of \mathbb{R}^n with a log-concave density in this subspace. Usually this affine subspace is \mathbb{R}^n itself.

For example, any Gaussian measure in \mathbb{R}^n is log-concave, because its density relative to the affine subspace where it is supported is of the form

$$c_A \exp(-\langle A(x-b), (x-b)\rangle)$$

for a symmetric, positive-definite operator A, a number $C_A>0$ and a vector $b\in\mathbb{R}^n$. The quadratic function $x\to \langle A(x-b),(x-b)\rangle$ is clearly convex, and hence the Gaussian measure is log-concave. The uniform probability measure on any bounded convex set, is log-concave as well. On the real line, it is very common to encounter log-concave distributions; pretty much, a typical distribution that decays exponentially or faster at infinity tends to be log-concave; exponential decay at infinity is indeed a necessary condition for log-concavity. Thus the exponential distribution on $[0,\infty)$ is log-concave, as well as beta and gamma distributions with certain parameters and the double-exponential probability density

$$\exp(-2|x|) \qquad (x \in \mathbb{R}).$$

Operations that preserve log-concavity include:

1. Linear images. If X is a log-concave random vector in \mathbb{R}^n , then for any subspace $E\subseteq\mathbb{R}^n$ also

$$Proj_E(X)$$

is log-concave, by Prekopa-Leindler. It follows that for any linear (or affine) map $T: \mathbb{R}^n \to \mathbb{R}^m$, the random vector T(X) is log-concave.

2. Pointwise product. If f_1, \ldots, f_N are log-concave functions, then so is the product $\prod_{i=1}^N f_i$. It follows that if a polynomial P has only real roots and is positive on an interval I, then its restriction to I is log-concave. Indeed,

$$P(x) = c \cdot 1_I(x) \cdot \prod_{i=1}^{N} (x - \lambda_i)$$

for some interval $I \subseteq \mathbb{R}$, a real number $c \in \mathbb{R}$ and $\lambda_1, \ldots, \lambda_N \in \mathbb{R} \setminus I$. Since $|x - \lambda_i|$ is log-concave on I, the same applies for P.

- 3. Convolution. If $f,g:\mathbb{R}^n\to [0,\infty)$ are log-concave, then f(y)g(x-y) is log-concave on $\mathbb{R}^n\times\mathbb{R}^n$, and consequently the same applies for its marginal f*g.
- 4. Weak limits. It is an exercise to deduce from the Prékopa-Leindler inequality that if $(\mu_N)_{N\geq 1}$ is a sequence of log-concave concave probability measures converging weakly to a probability measure μ , then μ is also log-concave. This is not an obvious fact; think of the case where μ_N tend to a measure supported on a lower dimensional subspace.

Proposition 6.3 ("How to think on 1D log-concave random variables"). Let $X \in \mathbb{R}$ be an isotropic, log-concave random variable, i.e., $\mathbb{E}X = 0$ and Var(X) = 1. Write ρ for the log-concave density of X. Then for all $x \in \mathbb{R}$,

$$c'1_{\{|x| < c''\}} \le \rho(x) \le Ce^{-c|x|},$$

where c, c', c'', C > 0 are universal constants.

Sketch of proof. For the upper bound, if $\rho(b) < \rho(a)/2$ for some a < b, then ρ decays exponentially and in fact $\rho(x) \le \rho(b)2^{-x/(b-a)}$ for all x > b. As for the lower bound, it is enough to show that $\rho(x) > c'$ for some x > c'' and for some x < -c''. It is an exercise to filling in the details.

Corollary 6.4 ("reverse Hölder inequalities", Berwald [1, 3]). For any isotropic, log-concave, real-valued random variable X and for any p > -1,

$$c \cdot \min\{p+1,1\} \le ||X||_p = (\mathbb{E}|X|^p)^{1/p} \le C(p+2),$$
 (6)

where c, C > 0 are universal constants.

The case p = 0 in (6) is interpreted by continuity, i.e.,

$$||X||_0 = \exp(\mathbb{E}\log|X|).$$

This is not a norm, yet a nice feature is its multiplicativity: for any random variables X and Y, possibly dependent,

$$||XY||_0 = ||X||_0 ||Y||_0.$$

Proof of Corollary 6.4. Begin with the inequality on the right-hand side. By the monotonicity of $p \mapsto ||X||_p$, it is enough to look at p > 0. In this case,

$$\|X\|_p^p = \int_{-\infty}^{\infty} |t|^p \rho(t) dt \le C \int_{-\infty}^{\infty} |t|^p e^{-c|t|} dt = \frac{2C}{c^{p+1}} \Gamma(p+1) \le (\tilde{C}p)^p,$$

where we used the fact that for integer p, we have $\Gamma(p+1) = p! \le p^p$. For the lower bound, by monotonicity it suffices to look at p < 0. Setting $q = -p \in (0,1)$ we have

$$\mathbb{E}\frac{1}{|X|^q} \leq C \int_{-\infty}^{\infty} \frac{1}{|t|^q} e^{-c|t|} dt \leq \frac{C'}{1-q}$$

and hence

$$||X||_p = \left(\mathbb{E}\frac{1}{|X|^q}\right)^{-1/q} \ge \left(C'(1-q)\right)^{1/q} \ge \tilde{C}(1-q).$$

For instance, we learn from Corollary 6.4 that when X is a centered, log-concave random vector in \mathbb{R}^n , then

$$\mathbb{E}\langle X, \theta \rangle^4 \le C \left(\mathbb{E}\langle X, \theta \rangle^2 \right)^2, \tag{7}$$

for a universal constant C>0 (in fact, C=9 is optimal here, see Eitan [6]). Indeed, if $\sigma=(\mathbb{E}\langle X,\theta\rangle^2)^{1/2}$ then $\langle X,\theta\rangle/\sigma$ is an isotropic, log-concave random variable, and (7) follows from Corollary 6.4.

Corollary 6.5 ("Reverse Hölder inequalities for polynomials"). Let X be a real-valued, log-concave random variable, and let $f : \mathbb{R} \to \mathbb{R}$ be a polynomial of degree at most d. Then for any 0 ,

$$||f(X)||_q \le C_{q,d} \cdot ||f(X)||_p$$

for some constant $C_{q,d}$ depending only on q and d.

Proof. Following Bobkov [2], we may assume that f is a monic polynomial in one real variable, hence

$$f(X) = \prod_{i=1}^{d} (X - z_i)$$

for some $z_1, \ldots, z_d \in \mathbb{C}$. Consequently, by Hölder's inequality and by Corollary 6.4,

$$||f(X)||_q = \left\| \prod_{i=1}^d (X - z_i) \right\|_q \le \prod_{i=1}^d ||X - z_i||_{dq}$$

$$\le \prod_{i=1}^d Cd(q+1)||X - z_i||_0 = (Cd(q+1))^d ||f(X)||_0.$$

We remark that Corollary 6.5 remains valid verbatim if one replaces "real-valued log-concave random variable" by "log-concave random vector in a finite-dimensional normed space"; see Bourgain [4].

Corollary 6.6 (Hensley [8], Fradelizi [9]). Let $K \subseteq \mathbb{R}^n$ be a centered convex body. Assume that the random vector X that is distributed uniformly in K, is isotropic (or more generally, that Cov(X) is a scalar matrix). Then for any hyperplanes $H_1, H_2 \subseteq \mathbb{R}^n$ passing through the origin,

$$Vol_{n-1}(K \cap H_1) \leq C \cdot Vol_{n-1}(K \cap H_2)$$

where C > 0 is a universal constant. In fact, $C \le \sqrt{6}$.

Proof. Let $\theta \in S^{n-1}$ and denote

$$\rho_{\theta}(t) = \frac{Vol_{n-1}(K \cap (t\theta + \theta^{\perp}))}{Vol_{n}(K)}.$$

Then ρ_{θ} is the density of the random variable $X \cdot \theta$, which is log-concave and isotropic. According to Proposition 6.3, for any $x \in \mathbb{R}$,

$$c'1_{\{|x| \le c''\}} \le \sigma \rho_{\theta}(x\sigma) \le Ce^{-c|x|}$$

In particular,

$$c \le \rho_{\theta}(0) \le C$$
,

for some universal constants c, C > 0. Thus, for $\theta_1, \theta_2 \in S^{n-1}$,

$$\frac{Vol_{n-1}(K \cap \theta_1^{\perp})}{Vol_{n-1}(K \cap \theta_2^{\perp})} = \frac{\rho_{\theta_1}(0)}{\rho_{\theta_2}(0)} \le \frac{C}{c} \le C'.$$

Thus, up a multiplicative universal constant, volumes of hyperplane sections of K are closely related to the covariance matrix of the uniform distribution of K. When K is not isotropic, we have the following:

Corollary/Exercise 6.7. Let $K \subseteq \mathbb{R}^n$ be a convex body of volume one and let $X \sim Unif(K)$. Then,

$$\sup_{H} Vol_{n-1}(K \cap H) \sim \sqrt{\|\operatorname{Cov}(X)^{-1}\|_{op}},$$

where the supremum runs over all hyperplanes $H \subseteq \mathbb{R}^n$, and where $||A||_{op}$ is the operator norm of the matrix $A \in \mathbb{R}^{n \times n}$.

A useful geometric parameter is concerned with the relation between the covariance matrix and entropy:

Definition 6.8. For a convex body $K \subseteq \mathbb{R}^n$ we define its isotropic constant to be

$$L_K = \left(\frac{\det \operatorname{Cov}(K)}{Vol_n(K)^2}\right)^{\frac{1}{2n}}$$

where Cov(K) is the covariance matrix of the uniform probability distribution on K. More generally, the isotropic constant of an absolutely continuous, log-concave random vector X in \mathbb{R}^n is

$$L_X = \left(\frac{\det \operatorname{Cov}(X)}{e^{2\operatorname{Ent}(X)}}\right)^{\frac{1}{2n}},\tag{8}$$

where $\operatorname{Ent}(X) = -\int_{\mathbb{R}^n} \rho \log \rho$ when ρ is the density of X.

The isotropic constant measures the relation between the (differential) entropy of a distribution, and the determinant of its covariance matrix. These two parameters measure the "size" of a distribution in \mathbb{R}^n , in a way which is *invariant* under the group of volume-preserving linear transformations of \mathbb{R}^n . Clearly, for any invertible affine transformation $T: \mathbb{R}^n \to \mathbb{R}^n$,

$$L_{T(X)} = L_X.$$

The entropy of a log-concave random vector in \mathbb{R}^n is a rather robust quantity, and up to an additive error of O(n) may be computed in many ways:

Lemma 6.9. Let X be a log-concave random vector in \mathbb{R}^n with density $\rho = e^{-\psi}$. Then,

$$\psi(\mathbb{E}X) \le \text{Ent}(X) \le \inf \psi + n.$$
 (9)

In fact, most of the mass of the random vector X is located in regions where ψ is at most $\inf \psi + 2n$. Indeed,

$$\mathbb{E}e^{\frac{\psi(X)}{2}} < e^{\frac{\inf\psi}{2} + (\ln 2)n}.$$

and hence by Markov-Chebyshev,

$$\mathbb{P}\left(\psi(X) \ge \inf \psi + 2n\right) \le e^{-cn},$$

for a universal constant c > 0.

Proof. By approximation, we may assume that ρ is smooth and positive in \mathbb{R}^n ; this will allow us to neglect boundary terms in integration by parts below. The function ψ is convex, and hence it lies above its tangents. That is, for any $x, y \in \mathbb{R}^n$,

$$\psi(y) \ge \psi(x) + \langle \nabla \psi(x), y - x \rangle. \tag{10}$$

By Jensen's inequality,

$$\psi(\mathbb{E}X) \le \mathbb{E}\psi(X) = \mathrm{Ent}(X),$$

and the left-hand side inequality in (9) is proven. For proving the right-hand side inequality, note that by (10), for any $y \in \mathbb{R}^n$,

$$\operatorname{Ent}(X) = \mathbb{E}\psi(X) \le \mathbb{E}\left[\psi(y) - \nabla\psi(X) \cdot (y - X)\right]$$
$$= \psi(y) + \mathbb{E}\left[\nabla\psi(X) \cdot (X - y)\right].$$

Now,

$$\mathbb{E}\langle\nabla\psi(X),X\rangle = \int_{\mathbb{R}^n} \left[\sum_{i=1}^n \partial^i \psi(x) x_i\right] e^{-\psi(x)} dx$$
$$= -\sum_{i=1}^n \int_{\mathbb{R}^n} x_i \partial^i \left(e^{-\psi(x)}\right) = \sum_{i=1}^n \int_{\mathbb{R}^n} e^{-\psi} = n, \qquad (11)$$

and similarly $\mathbb{E}\nabla\psi(X)\cdot y=0$. Inequality (9) thus follows by taking the infimum in (11) over all $y\in\mathbb{R}^n$. Additionally,

$$\mathbb{E}e^{\frac{\psi(X)}{2}} = e^{\frac{\psi(y)}{2}} \int_{\mathbb{R}^n} e^{-\frac{\psi(x) + \psi(y)}{2}} dx \le e^{\frac{\psi(y)}{2}} \int_{\mathbb{R}^n} e^{-\psi\left(\frac{x+y}{2}\right)} dx = 2^n e^{\frac{\psi(y)}{2}} \int e^{-\psi\left(\frac{x+y}{2}\right)} dx = 2^n e^{\frac{\psi(y)}{2}}.$$

The lemma follows by taking the infimum over all $y \in \mathbb{R}^n$.

Which probability measure in \mathbb{R}^n has the smallest isotropic constant?

Claim 6.10. For any absolutely-continuous, log-concave random vector X in \mathbb{R}^n ,

$$L_X \ge \frac{1}{\sqrt{2\pi e}}$$

with equality if and only if X is a Gaussian random vector.

Proof. Log-concavity is irrelevant for this claim, which basically asserts the well-known fact that the Gaussian distribution has maximal entropy among all distributions with a given covariance matrix.

Translating, we may assume that X is centered. Let us write ρ for the density of X. Then for any positive probability density $\gamma : \mathbb{R}^n \to (0, \infty)$,

$$\mathbb{E}\log\frac{\gamma(X)}{\rho(X)} \le \mathbb{E}\left[\frac{\gamma(X)}{\rho(X)} - 1\right] = \int_{\mathbb{R}^n} \frac{\gamma(x)}{\rho(x)} \rho(x) dx - 1 = 0.$$

In particular, if γ is the density of a centered Gaussian random vector G in \mathbb{R}^n with $\operatorname{Cov}(X) = \operatorname{Cov}(G)$, then $\log \gamma$ is a quadratic function plus a constant and

$$\operatorname{Ent}(X) = -\mathbb{E}\log\rho(X) \le -\mathbb{E}\log\gamma(X) = -\mathbb{E}\log\gamma(Z) = \operatorname{Ent}(G).$$

By the end of these lectures we will (hopefully) complete the proof that

$$L_{\rm Y} < C$$

for a universal constant C>0, due to Klartag and Lehec [11] following Guan [7]. The optimal value of C remains unknown. The symmetric case of the *strong slicing conjecture* suggests that when $K\subseteq\mathbb{R}^n$ is a centrally-symmetric convex body,

$$L_K \le L_{[0,1]^n} = \frac{1}{\sqrt{12}}.$$

If true, then the Minkowski lattice conjecture follows, see Magazinov [14] and references therein. The Minkowski lattice conjecture suggests that if $L\subseteq\mathbb{R}^n$ is a lattice of determinant one, then each of its translates intersects the set

$$\left\{ x \in \mathbb{R}^n : \prod_{i=1}^n |x_i| \le \frac{1}{2^n} \right\}.$$

This was proven in two dimensions by Minkowski in 1908. Moreover, the isotropic constant of any simplex $\Delta^n \subseteq \mathbb{R}^n$ equals

$$L_{\Delta^n} = \frac{(n!)^{1/n}}{(n+1)^{(n+1)/(2n)}\sqrt{n+2}} \approx \frac{1}{e}.$$

If the isotropic constant is maximized for the simplex among all convex bodies – this is the non-symmetric strong slicing conjecture – then the Mahler conjecture follows in the non-symmetric case. See [13]. This Mahler conjecture, which we discussed earlier, suggests that among all convex bodies $K \subseteq \mathbb{R}^n$, the volume product

$$Vol_n(K) \cdot Vol_n(K^{\circ})$$

is minimized when K is a centered simplex. This was proven in two dimensions by Mahler in 1938.

A useful tool for analyzing a log-concave random vector X in \mathbb{R}^n is its *logarithmic* Laplace transform, defined via

$$\Lambda(y) = \Lambda_X(y) = \log \mathbb{E}e^{\langle X, y \rangle} \qquad (y \in \mathbb{R}^n).$$

Proposition 6.11 ("The logarithmic Laplace transform of log-concave measures"). Let X be a log-concave random vector in \mathbb{R}^n . Then the following hold:

(a) The set $\Omega=\{y\in\mathbb{R}^n\,;\,\Lambda(y)<\infty\}$ is an open, convex set. The function

$$\Lambda:\Omega\to\mathbb{R}$$

is smooth and strongly-convex.

- (b) If X is absolutely-continuous, then the map $\nabla \Lambda$ is a diffeomorphism from Ω onto the convex hull of the interior of the support of X.
- (c) Suppose that X has density ρ , and for $y \in \Omega$ consider the "tilted probability measure"

$$\rho_y(x) = \rho(x)e^{x \cdot y - \Lambda(y)}.$$

Let X_y be a random vector with density ρ_y and denote $a_y = \mathbb{E}X_y$. Then, for any $y \in \Omega$,

$$\nabla \Lambda(y) = \mathbb{E}X_y, \qquad \nabla^2 \Lambda(y) = \operatorname{Cov}(X_y), \qquad \nabla^3 \Lambda(y) = \mathbb{E}(X_y - a_y)^{\otimes 3}.$$

Sketch of Proof. Assume that X is absolutely-continuous. If $y \in \Omega$ then ρ_y is an integrable, log-concave density. As such, it decays exponentially at infinity (see exercise below), and hence Ω contains a neighborhood of y. This shows that Ω is open, and that Λ is smooth. Conclusion (c) is a direct computation, which also implies the strong convexity of Λ and the fact that

$$\Omega \ni x \mapsto \nabla \psi(x) \in \mathbb{R}^n$$

is one-to-one and a diffeomorphism onto its image, which is an open set. Write K for the interior of the convex hull of the support of X. Then (c) implies that

$$\nabla \psi(\Omega) \subseteq K$$
,

since the barycenter of any absolutely-continuous distribution in K is contained in K. We leave the proof of the fact that $\nabla \psi(\Omega) = K$ to the reader. This fact will not be used below.

Exercises.

- 1. The Maurey-Pisier proof of Gaussian concentration.
 - (a) Let X and Y be two independent, standard Gaussian random vectors in \mathbb{R}^n . For $\theta \in [0, \pi/2]$ set

$$X_{\theta} = (\sin \theta)X + (\cos \theta)Y.$$

Prove that $(X_{\theta}, \partial X_{\theta}/\partial \theta)$ coincides in distribution with (X, Y).

(b) Let $F: \mathbb{R}^n \to \mathbb{R}$ be a locally-Lipschitz function and let $\varphi: \mathbb{R}^n \to \mathbb{R}$ be a convex function. Prove that

$$\mathbb{E}\varphi(F(X) - F(Y)) = \mathbb{E}\varphi\left(\int_0^{\pi/2} \left\langle \nabla F(X_\theta), \frac{\partial X_\theta}{\partial \theta} \right\rangle d\theta\right).$$

(c) Denote $E = \mathbb{E}F(X)$. Conclude that for any $\lambda > 0$,

$$\mathbb{E} e^{\lambda(F(X)-E)} \leq \mathbb{E} e^{\lambda\pi\langle\nabla F(X),Y\rangle/2} = \mathbb{E} e^{\lambda^2\pi^2|\nabla F(X)|^2/8}.$$

(d) Conclude that if F is 1-Lipschitz, then for all t > 0,

$$\mathbb{P}(|F(X) - E| > t) < 2e^{-2t^2/\pi^2}.$$

2. Let $\mu, \mu_1, \mu_2, \ldots$ be log-concave probability measures on \mathbb{R}^n . Assume that $\mu_N \longrightarrow \mu$ weakly, i.e., that for any continuous, compactly-supported function $\varphi: \mathbb{R}^n \to \mathbb{R}$,

$$\lim_{N\to\infty}\int_{\mathbb{R}^n}\varphi d\mu_N=\int_{\mathbb{R}^n}\varphi d\mu.$$

Suppose that μ_N is log-concave for all N. Prove that μ is log-concave.

- 3. Complete the proof of Proposition 6.3
- 4. Let $\rho: \mathbb{R}^n \to [0, \infty)$ be a log-concave probability density. Prove that there exist A, B > 0 such that for all $x \in \mathbb{R}^n$,

$$\rho(x) \le Ae^{-B|x|}.$$

5. Complete the proof of Proposition 6.11.

References

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Request. Please e-mail me at boaz.klartag@weizmann.ac.il with any comments, corrections, or suggestions regarding these lecture notes. In addition, if you are able to produce *tikz* figures similar to (or better than) those drawn on the blackboard in class, I would be glad to include them in the notes.